RESEARCH PAPER | Mathematics | Volume : 5 | Issue : 5 | May 2015 | ISSN - 2249-555X

Bayesian Nonparametric Regression Model Based on Spline

KEYWORDS: Nonparametric Regression , Spline , Bayes Approach , Distribution , Posterior Distribution , Bayes factors .

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ABSTRACT *In this paper , Bayesian approach to nonparametric regression model described . The nonparametric regression model is assumed to be a smooth spline . Bayes approach based on Markov chain Monte Carlo (MCMC) employed to make inferences on the resulting spline nonparametric model coefficients under some conditions on the prior and design matrix. We investigate the posterior density and identify the analytic form of the Bayes factors.*

1. Introduction | In the Bayesian approach to inference the fixed but unknown parameters are viewed as a random variables . It is well known that the Bayes estimate under squared error loss of any subvector of the parameters vector is the mean of its posterior distribution,[10],[12]. | Markov Chain Monte Carlo (MCMC) method depends on partition of difficult and compound models into simple ones which can be manipulated and easily analyzed , specially for the posterior distribution which are not easy to find their final formula. | The Bayesian, Bayesian nonparametric and Bayesian Semiparametric regression models , were studied by many researchers for example DeRobertis and Hartign in (1981) discussed the Bayesian inference using intervals of measeres,[7]. Berger in (1985) introduced the statistical decision theory and Bayesian analysis,[4] . Lenk in (1999) presented the Bayesian inference of a Semiparametric regression model using Fourier representation,[11]. Zhao in (2000) studied the Bayesian approach to the nonparametric function estimation problems such as nonparametric regression and signal estimation and he considered the asymptotic of Bayes procedures for conjugate (Gaussian) priors,[15]. Angers and Delampady in (2001) used the Bayesian approach to the nonparametric regression model using a wavelet basis and performed the subsequence estimations,[1]. Dass and Lee in (2002) presented a note on the consistency of Bayes factors for testing point null versus nonparametric alternative,[6].Ghosh, J.K. and Ramamoorthi, R.V. in (2003) studied Bayesian nonparametric,[9]. Angers and Delampady in (2004) studied fuzzy sets in hierarchical Bayes and robust Bayes inference,[2]. Ghosh, J.K. , Delampady M. and samanta, T. in (2006) presented Bayesian analysis and discussed the theory and methods,[8]. Angers and Delampady in (2008) discussed fuzzy sets in nonparametric Bayes regression by using wavelet and membership functions and they treated the membership functions as likelihood functions for the model,[3] . Choi , Lee and Roy in (2008) investigated the large sample property of the Bayes factor for testing the parametric null model against the Semiparametric alternative model,[5] Under some conditions on the prior and design matrix , and using algebraic smoothing , they identified the analytic form of the Bayes factor and showed that the Bayes factor is consistent. Osaba and Mitaim in (2011) examined Bayesian with adaptive fuzzy priors and the likelihoods member,[13]. Pelenis in (2012) studied the Bayesian Semiparametric regression and considered a Bayesian estimation of restricted conditional moment models with linear regression as a particular example,[14]. | This paper came to shed light on the nonparametric regression model which has a nonparametric function is assumed to be a smooth spline , as well as the error term which has normal distribution with mean zero and variance σ_ε^2. | In this paper , Bayesian approach based on Markov chain Monte Carlo (MCMC) employed to make inferences on the resulting spline nonparametric regression model coefficients under some conditions on the prior distribution and design matrix. | We investigate the posterior density and identify the analytic form of the Bayes factors to choose between a fully Bayesian spline nonparametric regression model with $(p+k+1)$ of parameters against a Bayesian spline nonparametric regression model with $(p+q+1)$ of parameters , where $q<k$, | 2. Description of the problem | Consider the model: |

$$y_i = m(x_i) + \varepsilon_i \quad , \quad i = 1,2,...,n \quad . \qquad (1)$$

| Where the unobserved errors $\varepsilon_1, \varepsilon_2, ..., \varepsilon_n$ are known to be i.i.d. normal with mean zero and covariance $\sigma_\varepsilon^2 I_n$ with σ_ε^2 unknown and is nonparametric component . By using spline of

$$y_i = \beta_0 + \beta_1 x_i + ... + \beta_p x_i^p + \sum_{j=1}^{k} \beta_{j+p}(x_i - t_j)_+^p + \varepsilon_i, \qquad (2)$$

degree get : | where $t_1, t_2, ..., t_k$ are inner knots . The model (2) is rewritten as follows: |

$$Y = X\beta + \varepsilon, \qquad (3)$$

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$$Y=\begin{bmatrix} y_1\\ \vdots \\ y_n \end{bmatrix}_{n\times 1},\quad \boldsymbol{\beta}=\begin{bmatrix} \beta_0\\ \beta_1\\ \vdots\\ \beta_p\\ \beta_{p+1}\\ \vdots\\ \beta_{p+k+1} \end{bmatrix},\quad \boldsymbol{\varepsilon}=\begin{bmatrix} \varepsilon_1\\ \vdots\\ \varepsilon_n \end{bmatrix},\quad \boldsymbol{\varepsilon}\overset{iid}{\sim} N(0,\sigma_\varepsilon^2 I),$$

where

$$X=\begin{bmatrix} 1 & x_1 & \cdots & x_1^p & (x_1-t_1)_+^p & \cdots & (x_1-t_k)_+^p \\ 1 & x_2 & \cdots & x_2^p & (x_2-t_1)_+^p & \cdots & (x_2-t_k)_+^p \\ \vdots & \vdots & \cdots & \vdots & \vdots & \cdots & \vdots \\ 1 & x_n & \cdots & x_n^p & (x_n-t_1)_+^p & \cdots & (x_n-t_k)_+^p \end{bmatrix}_{n\times(p+k+1)}$$

| The estimation of the parameters β entails minimizing the spline least squares criterion : |

$$\|Y - X\beta\|^2. \tag{4}$$

The least squares estimators from (4) are :

$$\hat{\beta} = (X^T X)^{-1} X^T Y, \tag{5}$$

and the fitted valued are:

$\hat{Y} = X\hat{\beta} = HY$, where H is the smoothing matrix given by :

$$H = X(X^T X)^{-1} X^T.$$

3. The Prior Distribution

To specify a complete Bayesian model , we need a prior distribution on $(\beta,\sigma_\varepsilon^2)$. If a proper prior is desired , one could use a $N(0,\sigma_\beta^2 I)$ prior with σ_β^2 so large that for all intents and purposes , the normal distribution is uniform on the range of β . Therefore , we will use $\pi_0(\beta)\equiv 1$. As well as we will assume that the prior on σ_ε^2 is inverse gamma with parameters α_ε and β_ε i.e.

$$\pi_0(\sigma_\varepsilon^2) = \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)}(\sigma_\varepsilon^2)^{-(\alpha_\varepsilon+1)}\exp\left(-\frac{\beta_\varepsilon}{\sigma_\varepsilon^2}\right), \tag{6}$$

where α_ε and β_ε are hyperparameters that determine the priors and must be chosen by the statistician .

4. Posterior Distribution

From the model (3) we have

$$Y \mid \beta,\sigma_\varepsilon^2 \sim N(X\beta,\sigma_\varepsilon^2 I_n). \tag{7}$$

Then the likelihood function $L(Y\mid\beta,\sigma_\varepsilon^2)$ can be expressed as:

$$L(Y\mid\beta,\sigma_\varepsilon^2) = \frac{1}{(2\pi\sigma_\varepsilon^2)^{n/2}}\exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(Y-X\beta)^T(Y-X\beta)\right\}$$

$$= \frac{1}{(2\pi\sigma_\varepsilon^2)^{n/2}} \exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(Y - X\hat{\beta})^T (Y - X\hat{\beta})\right\} \times$$

$$\exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(\beta - \hat{\beta})^T X^T X(\beta - \hat{\beta})\right\}. \quad (8)$$

Then the joint posterior density of the coefficients β and the error variance σ_ε^2 given by the expression

$$\pi_1(\beta, \sigma_\varepsilon^2 \mid Y) \propto L(Y \mid \beta, \sigma_\varepsilon^2)\pi_0(\beta, \sigma_\varepsilon^2) \quad (9)$$

$$\propto (\sigma_\varepsilon^2)^{-n/2} \exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(Y - X\hat{\beta})^T (Y - X\hat{\beta})\right\}$$

$$\times \exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(\beta - \hat{\beta})^T X^T X(\beta - \hat{\beta})\right\} \times \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)}(\sigma_\varepsilon^2)^{-(\alpha_\varepsilon+1)} \exp\left\{-\frac{\beta_\varepsilon}{\sigma_\varepsilon^2}\right\}$$

$$\propto (\sigma_\varepsilon^2)^{-(n/2+\alpha_\varepsilon+1)} \exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(\beta - \hat{\beta})^T X^T X(\beta - \hat{\beta})\right\} \times$$

$$\exp-\left\{\frac{\frac{1}{2}(Y - X\hat{\beta})^T (Y - X\hat{\beta}) + \beta_\varepsilon}{\sigma_\varepsilon^2}\right\}. \quad (10)$$

From this expression , we deduce the following conditional and marginal posterior distributions

$$\pi_1(\beta \mid \sigma_\varepsilon^2, Y) \propto \exp\left\{-\frac{1}{2\sigma_\varepsilon^2}(\beta - \hat{\beta})^T X^T X(\beta - \hat{\beta})\right\}, \quad (11)$$

and

$$\pi_1(\sigma_\varepsilon^2 \mid \beta, Y) \propto (\sigma_\varepsilon^2)^{-(n/2+\alpha_\varepsilon+1)} \exp-\left\{\frac{\frac{1}{2}(Y - X\hat{\beta})^T (Y - X\hat{\beta}) + \beta_\varepsilon}{\sigma_\varepsilon^2}\right\}. \quad (12)$$

Therefore , it follows that

$$\beta \mid \sigma_\varepsilon^2, Y \sim N(\hat{\beta}, \sigma_\varepsilon^2 (X^T X)^{-1}) \quad (13)$$

$$\sigma_\varepsilon^2 \mid \beta, Y \sim IG\left(\alpha_\varepsilon + \frac{n}{2}, \beta_\varepsilon + \frac{1}{2}(Y - X\hat{\beta})^T (Y - X\hat{\beta})\right) \quad (14)$$

5. Model checking and Bayes factors

We would like to choose between a fully Bayesian spline nonparametric regression model with $(p+k+1)$ of parameters against a Bayesian spline nonparametric regression model with $(p+q+1)$ of parameters , where $q < k$, by using Bayes factors for two hypotheses

$$\left.\begin{array}{l} H_0 : y_i = \boldsymbol{\beta}_0 + \boldsymbol{\beta}_1 x_i + \ldots + \boldsymbol{\beta}_p x_i^p + \sum_{j=1}^{q} \boldsymbol{\beta}_{j+p} (x_i - t_j)_+^p + \boldsymbol{\varepsilon}_i \quad or \quad Y = X^0 \boldsymbol{\beta}^0 + \boldsymbol{\varepsilon} \\ vesus \\ H_1 : y_i = \boldsymbol{\beta}_0 + \boldsymbol{\beta}_1 x_i + \ldots + \boldsymbol{\beta}_p x_i^p + \sum_{j=1}^{k} \boldsymbol{\beta}_{j+p} (x_i - t_j)_+^p + \boldsymbol{\varepsilon}_i \quad or \quad Y = X\boldsymbol{\beta} + \boldsymbol{\varepsilon} \end{array}\right\}, \tag{15}$$

where $\boldsymbol{\beta}^0$ is $(p+q+1)\times 1$ vectors of parameters , X^0 is an $n\times(p+q+1)$ design matrix and $q<k$. We compute the Bayes factor , B_{01} , of H_0 relative to H_1 for testing problem (15) as follows

$$B_{01}(Y) = \frac{m(Y \mid H_0)}{m(Y \mid H_1)}, \tag{16}$$

where $m(Y \mid H_0)$ is the marginal density of Y under model $H_i, i = 0,1$.

We have:

$$m(Y \mid H_0) = \int \left(\int f(Y \mid \boldsymbol{\beta}^0, \boldsymbol{\sigma}_\varepsilon^2) \boldsymbol{\pi}_1(\boldsymbol{\beta}^0 \mid \boldsymbol{\sigma}_\varepsilon^2) \boldsymbol{\pi}_0(\boldsymbol{\sigma}_\varepsilon^2) d\boldsymbol{\beta}^0 \right) d\boldsymbol{\sigma}_\varepsilon^2$$

$$= (2\boldsymbol{\pi})^{-\frac{n}{2}} \frac{\boldsymbol{\beta}_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\boldsymbol{\alpha}_\varepsilon)} \int (\boldsymbol{\sigma}_\varepsilon^2)^{-\left(\frac{n}{2}+\boldsymbol{\alpha}_\varepsilon+1\right)} \exp\left\{ -\frac{\left(\frac{1}{2}\right)\left(Y - X^0\boldsymbol{\beta}^0\right)^T \left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon}{\boldsymbol{\sigma}_\varepsilon^2} \right\} d\boldsymbol{\sigma}_\varepsilon^2$$

$$= (2\boldsymbol{\pi})^{-\frac{n}{2}} \frac{\boldsymbol{\beta}_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\boldsymbol{\alpha}_\varepsilon)} \int (\boldsymbol{\sigma}_\varepsilon^2)^{-\left(\frac{n}{2}+\boldsymbol{\alpha}_\varepsilon+1\right)} \left(\frac{1}{2}\left(Y - X^0\boldsymbol{\beta}^0\right)^T \left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon \right)^{\left(\frac{n}{2}+\boldsymbol{\alpha}_\varepsilon+1\right)}$$

$$\times \left(\frac{1}{2}\left(Y - X^0\boldsymbol{\beta}^0\right)^T \left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon \right)^{-\left(\frac{n}{2}+\boldsymbol{\alpha}_\varepsilon+1\right)} \times$$

$$\exp\left\{ -\frac{\left(\frac{1}{2}\right)\left(Y - X^0\boldsymbol{\beta}^0\right)^T \left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon}{\boldsymbol{\sigma}_\varepsilon^2} \right\} d\boldsymbol{\sigma}_\varepsilon^2$$

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$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_{\varepsilon}^{\alpha_{\varepsilon}}}{\Gamma(\alpha_{\varepsilon})} \int \frac{\left(\frac{1}{2}\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}\right)^{\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)}}{(\sigma_{\varepsilon}^2)^{\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)}}$$

$$\times \exp\left\{-\frac{\left(\frac{1}{2}\right)\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}}{\sigma_{\varepsilon}^2}\right\}$$

$$\times\left(\frac{1}{2}\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}\right)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)} d\sigma_{\varepsilon}^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_{\varepsilon}^{\alpha_{\varepsilon}}}{\Gamma(\alpha_{\varepsilon})} \int \left(\frac{\left(\frac{1}{2}\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}\right)}{(\sigma_{\varepsilon}^2)}\right)^{\left(\frac{n}{2}+\alpha_{\varepsilon}+2\right)-1}$$

$$\times \exp\left\{-\frac{\left(\frac{1}{2}\right)\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}}{\sigma_{\varepsilon}^2}\right\}$$

$$\times\left(\frac{1}{2}\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}\right)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)} d\sigma_{\varepsilon}^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_{\varepsilon}^{\alpha_{\varepsilon}}}{\Gamma(\alpha_{\varepsilon})} \Gamma\left(\frac{n}{2} + \alpha_{\varepsilon} + 2\right)\left(\frac{1}{2}\left(Y - X^0\beta^0\right)^T\left(Y - X^0\beta^0\right) + \beta_{\varepsilon}\right)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)}, \quad (17)$$

and

$$m(Y \mid H_1) = \int\left(\int f(Y \mid \beta, \sigma_{\varepsilon}^2)\pi_1(\beta \mid \sigma_{\varepsilon}^2)\pi_0(\sigma_{\varepsilon}^2)d\beta\right)d\sigma_{\varepsilon}^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_{\varepsilon}^{\alpha_{\varepsilon}}}{\Gamma(\alpha_{\varepsilon})} \int (\sigma_{\varepsilon}^2)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)} \exp\left\{-\frac{\left(\frac{1}{2}\right)\left(Y - X\beta\right)^T\left(Y - X\beta\right) + \beta_{\varepsilon}}{\sigma_{\varepsilon}^2}\right\} d\sigma_{\varepsilon}^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)} \int (\sigma_\varepsilon^2)^{-\left(\frac{n}{2}+\alpha_\varepsilon+1\right)} \left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{\left(\frac{n}{2}+\alpha_\varepsilon+1\right)}$$

$$\times \left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{-\left(\frac{n}{2}+\alpha_\varepsilon+1\right)} \times$$

$$\exp\left\{-\frac{\left(\frac{1}{2}\right)(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon}{\sigma_\varepsilon^2}\right\} d\sigma_\varepsilon^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)} \int \frac{\left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{\left(\frac{n}{2}+\alpha_\varepsilon+1\right)}}{(\sigma_\varepsilon^2)^{\left(\frac{n}{2}+\alpha_\varepsilon+1\right)}}$$

$$\times \exp\left\{-\frac{\left(\frac{1}{2}\right)(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon}{\sigma_\varepsilon^2}\right\}$$

$$\times \left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{-\left(\frac{n}{2}+\alpha_\varepsilon+1\right)} d\sigma_\varepsilon^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)} \int \left(\frac{\left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)}{(\sigma_\varepsilon^2)}\right)^{\left(\frac{n}{2}+\alpha_\varepsilon+2\right)-1}$$

$$\times \exp\left\{-\frac{\left(\frac{1}{2}\right)(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon}{\sigma_\varepsilon^2}\right\}$$

$$\times \left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{-\left(\frac{n}{2}+\alpha_\varepsilon+1\right)} d\sigma_\varepsilon^2$$

$$= (2\pi)^{-\frac{n}{2}} \frac{\beta_\varepsilon^{\alpha_\varepsilon}}{\Gamma(\alpha_\varepsilon)} \Gamma\left(\frac{n}{2} + \alpha_\varepsilon + 2\right) \left(\frac{1}{2}(Y - X\beta)^T (Y - X\beta) + \beta_\varepsilon\right)^{-\left(\frac{n}{2}+\alpha_\varepsilon+1\right)}, \quad (18)$$

using the above derivations , the Bayes factor for testing problem (15) is then given by:

$$B_{01}(Y)=\frac{\left(\frac{1}{2}\left(Y-X^{0}\boldsymbol{\beta}^{0}\right)^{T}\left(Y-X^{0}\boldsymbol{\beta}^{0}\right)+\boldsymbol{\beta}_{\varepsilon}\right)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)}}{\left(\frac{1}{2}\left(Y-X\boldsymbol{\beta}\right)^{T}\left(Y-X\boldsymbol{\beta}\right)+\boldsymbol{\beta}_{\varepsilon}\right)^{-\left(\frac{n}{2}+\alpha_{\varepsilon}+1\right)}} \tag{19}$$

6. Simulation Results

In this section , we illustrate the effectiveness of our methodology , we generated observations from the model (1) with the following regression functions :

(i) $y_1=\exp(2\pi x)$,

(ii) $y_2=Sin(2\pi x)+(x-0.5)^2$.

The observations for x are generated from uniform distribution on the interval [0,1] . The sample size taken are $n=25,\ 50,\ 100,\ 150,\ 200$.

The goodness of fit of the estimated models quantified by computing the criterions average mean squared error $(AMSE)$ and average mean absolute error $(AMAE)$ which are defined as:

$$AMSE=\frac{1}{N}\sum_{i=1}^{N}MSE(x_i), \tag{20}$$

$$AMAE=\frac{1}{N}\sum_{i=1}^{N}MAE(x_i), \tag{21}$$

where MSE and MAE are mean squared error and mean absolute error criterions respectively.

Table(1) presents summary values of the $(AMSE)$ and $(AMAE)$ for the estimation method . From this table we can see that the values of $(AMSE)$ and $(AMAE)$ when $(n=200)$ are smaller than their values for the first test function , which were (0.0006407081) and (0.000175353) respectively. While the values of $(AMSE)$ and $(AMAE)$ are smaller when $(n=150)$ for the second test function were (0.0001740030) and (0.000454008)

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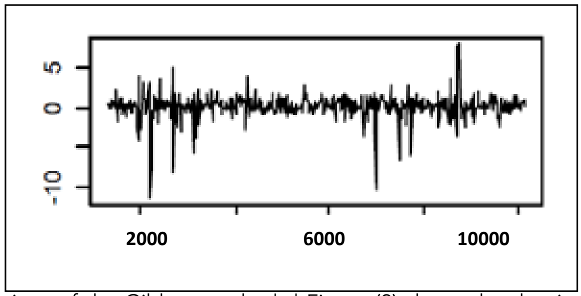
respectively. Figure (1) below shows the number for iterations of Gibbs sampler which used in this paper , which was (10000) iterations for this data . While figure (2) shows density estimates based on (10000) iterations of . σ_ε^2 | Table(1) results of the and criterions for Bayesian nonparametric regression model | | Figure (1) shows (10000) iterations of the Gibbs sampler | | Figure (2) shows the density estimates based on (10000) iterations of σ_ε^2 | |

Test functions	Sample size	$B_{01}(y)$	Evidence
y_1	25	2.127525×10^{-3}	Strongly favors H_1
	50	2.039452×10^{-9}	Strongly favors H_1
	100	3.776421×10^{-11}	Strongly favors H_1
	150	1.078844×10^{-14}	Strongly favors H_1
	200	4.883221×10^{-19}	Strongly favors H_1
y_2	25	6.543244×10^{-8}	Strongly favors H_1
	50	7.997665×10^{-19}	Strongly favors H_1
	100	4.988876×10^{-33}	Strongly favors H_1
	150	9.075544×10^{-40}	Strongly favors H_1
	200	3.111275×10^{-47}	Strongly favors H_1



7. Conclusions

The conclusions obtained throughout this paper are as follows:

(1)The posterior of $\boldsymbol{\beta}$ and $\boldsymbol{\sigma}_\varepsilon^2$ are respectively:

$$\boldsymbol{\beta} \mid \boldsymbol{\sigma}_\varepsilon^2, Y \sim N(\hat{\boldsymbol{\beta}}, \boldsymbol{\sigma}_\varepsilon^2 \quad \text{and} \quad \boldsymbol{\sigma}_\varepsilon^2 \mid \boldsymbol{\beta}, Y \sim IG\left(\boldsymbol{\alpha}_\varepsilon + \frac{n}{2}, \boldsymbol{\beta}_\varepsilon + \frac{1}{2}(Y - X\hat{\boldsymbol{\beta}})^T (Y - X\hat{\boldsymbol{\beta}})\right)$$

(2) The marginal density of Y under model $H_i, i = 0,1$ are :

$$m(Y \mid H_0) = (2\pi)^{-\frac{n}{2}} \frac{\boldsymbol{\beta}_\varepsilon^{\boldsymbol{\alpha}_\varepsilon}}{\Gamma(\boldsymbol{\alpha}_\varepsilon)} \Gamma\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 2\right)\left(\frac{1}{2}\left(Y - X^0\boldsymbol{\beta}^0\right)^T\left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon\right)^{-\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 1\right)},$$

and

$$m(Y \mid H_1) = (2\pi)^{-\frac{n}{2}} \frac{\boldsymbol{\beta}_\varepsilon^{\boldsymbol{\alpha}_\varepsilon}}{\Gamma(\boldsymbol{\alpha}_\varepsilon)} \Gamma\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 2\right)\left(\frac{1}{2}\left(Y - X\boldsymbol{\beta}\right)^T\left(Y - X\boldsymbol{\beta}\right) + \boldsymbol{\beta}_\varepsilon\right)^{-\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 1\right)}.$$

(3) The Bayes factor for testing problem (15) is given by the following form:

$$B_{01}(Y) = \frac{\left(\frac{1}{2}\left(Y - X^0\boldsymbol{\beta}^0\right)^T\left(Y - X^0\boldsymbol{\beta}^0\right) + \boldsymbol{\beta}_\varepsilon\right)^{-\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 1\right)}}{\left(\frac{1}{2}\left(Y - X\boldsymbol{\beta}\right)^T\left(Y - X\boldsymbol{\beta}\right) + \boldsymbol{\beta}_\varepsilon\right)^{-\left(\frac{n}{2} + \boldsymbol{\alpha}_\varepsilon + 1\right)}}$$

(4) In the simulation results , we concluded the following:

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(a) The values of and when are smaller thantheir values for the first test function , which were (0.0006407081) and (0.000175353) respectively. | (b) The values of and are smaller when for the second test function were (0.0001740030) and (0.000454008) respectively. | (c) The model corresponding to the first test function obtains the largest Bayes factor when followed by that the second test function when . | (d) The Bayes factor favors with strong evidence with all samples sizes for two test functions. | |

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