



## Analysis of Variance Application with Generalized Inverse Matrix

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**ABSTRACT**

In this study, a linear model of  $Y=X\beta+\varepsilon$  is used. The  $X'X$  matrix of the normal equations of  $(X'X)^{-1}X'Y$  obtained from this model is a non-full rank matrix. It is solved with generalized G-inverse since it is not the inverse of this non-full-rank matrix. The same results were obtained when the results of the generalized inverse matrix and variance analysis are compared. In non-full-rank models, solution with generalized inverse matrix is important in terms of implementation.

**KEYWORDS**

Non-full rank matrix, generalized inverse, linear model

### Introduction

Analysis of variance is examined while using linear models for the statistics of qualitative variables (Ipek, 1980). Extensions of these ideas to general operators have been made (Tseng, 1949). But no systematic study of the subject was made until 1955 when (Penrose, 1955), unannounced of the earlier work, redefined the Moore inverse in a different way.

In 1920, Moore published the first work on generalized inverses (Moore, 1920). Bott and Duffin (1953) defined what is called a constrained inverse of a square matrix, which is different from a G inverse and is useful in some applications. Chernoff (1953) considered an inverse of a singular nonnegative definite matrix, which is also not a G inverse but is useful in discussing some estimation problems. Mazmanoğlu and Kahramaner (2004) examined by generalized inverse matrices solution of qualitative variable non-full rank models. Nowadays, matrix algebra is used in all branches of mathematics and the sciences and constitutes the basis of most statistical procedures (Abdi and Williams, 2010).

The goal of this research, analysis of variance is carried out by generalized inverse matrix and present the application.

### Material and Method

The marks of students entered to postgraduate foreign language exam, in Ankara University Turkish and Foreign Language Research and Application Center, were examined. Whether or significant differences or not among marks of students whom of refer to Graduate School of Natural and Applied Sciences, Graduate School of Health and Applied Sciences, Graduate School of Social and Applied Sciences, Graduate School of Education and Applied Sciences among in the exam was tested. The data were used Senol Celik's (2006) in master thesis.

The model it endeavor is

$$Y = X\beta + \varepsilon$$

where  $Y$  is an  $N \times 1$  vector observations  $Y_i$ ,  $\beta$  is a  $p \times 1$  vector of parameters,  $X$  is an  $N \times p$  matrix of known values (in most cases 0 and 1) and  $\varepsilon$  is a vector of random error terms (Searle, 1997). Firstly,  $\varepsilon$  can be considered defined as  $\varepsilon = Y - E(Y)$  so that  $E(\varepsilon) = 0$  and  $E(Y) = X\beta$ .

Every element in  $\varepsilon$  is assumed to have variance  $\sigma^2$  and zero covariance with every other element.

$$V(\varepsilon) = E(\varepsilon\varepsilon') = \sigma^2 I_N \text{ (Rao, 1973; Searle, 1997).}$$

Thus

$$\varepsilon \sim N(0, \sigma^2 I) \text{ and } Y \sim N(X\beta, \sigma^2 I)$$

with normality being introduced subsequently (Searle, 1997). The normal equations corresponding to the model  $Y = X\beta + \varepsilon$  can be derived by least squares (Equation 1).

$$X'X\hat{\beta} = X'Y \tag{1}$$

Linear model of one-way ANOVA are examined.

$$y_{ij} = \mu + \alpha_i + \varepsilon_{ij}, \quad j = 1, 2, \dots, n_i, \quad i = 1, 2, \dots, k \tag{2}$$

where  $\mu$ : population mean,  $\alpha_i$ : the effect of type  $i$ ,  $\varepsilon_{ij}$ : random error term peculiar to the observation  $y_{ij}$  (Searle, 1997).

$$E(y_{ij}) = \mu + \alpha_i, \quad j=1,2,\dots,n_i, \quad i=1,2,\dots,k, \tag{3}$$

$$E(\varepsilon_i) = 0, V(\varepsilon_i) = \sigma^2 \text{ and } \varepsilon_i\text{'s independent (Akdeniz and Öztürk, 1996).}$$

They are easily rewritten in the form  $Y = X\beta + \varepsilon$  as

$$\begin{bmatrix} y_{11} \\ y_{12} \\ \vdots \\ \vdots \\ \vdots \\ y_{ij} \end{bmatrix} = \begin{bmatrix} 1 & 1 & 0 & \cdots & \cdots \\ 1 & 1 & 0 & \cdots & 0 \\ \vdots & 0 & 1 & 0 & \vdots \\ \vdots & \vdots & \vdots & 1 & \vdots \\ \vdots & \vdots & 0 & \vdots & \vdots \\ 1 & 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_i \end{bmatrix} + \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{12} \\ \vdots \\ \vdots \\ \vdots \\ \varepsilon_{ij} \end{bmatrix} \tag{4}$$

where,  $Y$ : vector of observation,  $X$  is the design matrix of 0 and 1,  $\beta$ : the vector of parameters,  $\varepsilon$ : the vector of error terms (Akdeniz and Öztürk, 1996; Searle 1997).  $Y = X\beta + \varepsilon$  model parameters are given in Table 1.

Table 1. Parameters of  $Y = X\beta + \varepsilon$  model

	Parameters					
Observations	$\mu$	$\alpha_1$	$\alpha_2$	$\alpha_3$	...	$\alpha_i$
$y_{11}$	1	1	0	0	$\vdots$	0
$y_{12}$	1	1	0	$\vdots$	$\vdots$	0
$\vdots$	$\vdots$	0	1	$\vdots$	$\vdots$	$\vdots$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$
$\vdots$	$\vdots$	$\vdots$	0	1	$\vdots$	$\vdots$
$y_{ij}$	1	$\vdots$	$\vdots$	$\vdots$	$\vdots$	1

In Table 1, as in equation (4), it is clear that the sum the last  $i$  columns equals the first column. The first column of  $X$  is all 1's; and every  $Y$  also contains just one  $\alpha$  and so the sum the last  $i$  columns is also 1's. Thus  $X$  not of full column rank. The consider the normal equations (1).  $XX'$  is square and symmetric (Equation 5).  $XX'$  is not of full column rank.

$$XX' = \begin{bmatrix} n & n_1 & n_2 & \cdots & n_i \\ n_1 & n_1 & \vdots & \vdots & 0 \\ n_2 & 0 & n_2 & \vdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ n_i & 0 & \cdots & \cdots & n_i \end{bmatrix} \tag{5}$$

Normal equations involve the vector  $X'Y$ , its elements are the inner products of the columns of  $X$  with the vector  $Y$ . The elements of  $X'Y$  are the sums of elements of  $Y$  from (4).

$$X'Y = \begin{bmatrix} 1 & 1 & \dots & \dots & \dots & 1 \\ 1 & 1 & 0 & \dots & \dots & 0 \\ 0 & \dots & 1 & \dots & 0 & \vdots \\ \vdots & \vdots & \vdots & 1 & \vdots & \vdots \\ 0 & \dots & \dots & \dots & 1 & 1 \end{bmatrix} \begin{bmatrix} Y_{11} \\ Y_{12} \\ \vdots \\ \vdots \\ Y_{ij} \end{bmatrix} = \begin{bmatrix} Y_{11} + Y_{12} + \dots + Y_{ij} \\ Y_{11} + \dots + Y_{1j} \\ Y_{21} + \dots + Y_{2j} \\ \vdots \\ Y_{ij} \end{bmatrix} = \begin{bmatrix} Y \\ Y_1 \\ Y_2 \\ \vdots \\ Y_i \end{bmatrix} \tag{6}$$

This is nature of X'Y in linear models a vector various subtotals o the Y observations (Equation 6). XX' is not of full rank, as in (5), the normal equations (1) cannot be solved with one solution  $\hat{\beta} = (X'X)^{-1}X'Y$ . To underline this it is written the normal equations as

$$X'X\beta^0 = X'Y \tag{7}$$

using the symbol to distinguish the many solutions of (7) from the solution that exist when X'X has full rank.  $\beta^0$  is used a solution GX'Y to (7), where G is a generalized inverse of X'X.

$$\begin{bmatrix} n & n_1 & n_2 & \dots & n_i \\ n_1 & n_1 & \vdots & \vdots & 0 \\ n_2 & 0 & n_2 & \vdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ n_i & 0 & \dots & \dots & n_i \end{bmatrix} \begin{bmatrix} \mu^0 \\ \alpha_1^0 \\ \alpha_2^0 \\ \vdots \\ \alpha_i^0 \end{bmatrix} = \begin{bmatrix} Y \\ Y_1 \\ Y_2 \\ \vdots \\ Y_{ij} \end{bmatrix} \tag{8}$$

X'X does not have full column and X'X has no inverse. Therefore, the normal equations (7) have no unique solution. To get any one of them it is found any generalized inverse G of X'X and write the corresponding solution as

$$\beta^0 = GX'Y \tag{9}$$

The notation  $\beta^0$  in equation (9) for a solution to the normal equations (7) emphasizes that what is derived by solving (7) is only a solution to the equations (Searle, 1997).

Let A be a matrix of order mxn. Then, generalized inverse (G-inverse) any nxm matrix A, denoted by  $A^- = G$ , is a matrix of order nxm that satisfies the relation AGA=A (Khuri, 2003; Kabe and Gupta, 2007).

**Results**

Whether or significant differences or not among marks of students whom of refer to postgraduate exam was investigated. Exam marks of total 380 students whom of refer to Graduate School of Natural and Applied Sciences, Graduate School of Health Sciences, Graduate School of Social Sciences, Graduate School of Education Sciences was researched. This problem is solved according to One-Way Analysis of Variance (ANOVA) It is reached to solution by Generalized Inverse. The model for analysis is the fixed-effects and the null hypothesis is (Chiang, 2003).

$$H_0 : \mu_1 = \mu_2 = \mu_3 = \mu_4 = 0$$

and

$H_1$  : (At least two institutions of student averages achievement are different).

$$(\mu_i \neq \mu_j) \quad i, j = 1, 2, 3, 4$$

Where,

$\mu_1$  = Institute of Science and Technology their students average achievement from foreign language exam

$\mu_2$  = Institute of Health Sciences their students average achievement from foreign language exam

$\mu_3$  = Institute of Social Sciences their students average achievement from foreign language exam

$\mu_4$  = Institute of Education Sciences their students average achievement from foreign language exam.

$$n_1 = n_2 = n_3 = n_4 = 95$$

$$n = n_1 + n_2 + n_3 + n_4 = 380 \text{ (Total observation)}$$

$$k = 4 \text{ (Number of group).}$$

Table 2. The total and average number of foreign language exam grade.

Institute	Total success notes	The average success
Science and Technology	$y_1 = \sum_{i=1}^{n_1} y_{1j} = \sum_{j=1}^{95} y_{1j} = 5398$	$\bar{y}_1 = \frac{\sum_{j=1}^{n_1} y_{1j}}{n_1} = \frac{5398}{95} = 56,821$
Health Sciences	$y_2 = \sum_{j=1}^{n_2} y_{2j} = \sum_{i=1}^{95} y_{2j} = 5748$	$\bar{y}_2 = \frac{\sum_{j=1}^{n_2} y_{2j}}{n_2} = \frac{5748}{95} = 60.505$
Social Science	$y_3 = \sum_{j=1}^{n_3} y_{3j} = \sum_{j=1}^{95} y_{3j} = 6001$	$\bar{y}_3 = \frac{y_3}{n_3} = \frac{6001}{95} = 63.168$
Education Science	$y_4 = \sum_{i=1}^{n_4} y_{4j} = \sum_{j=1}^{95} y_{4j} = 5673$	$\bar{y}_4 = \frac{y_4}{n_4} = \frac{5673}{95} = 59.716$
Overall Success	$Y = \sum_{i=1}^k y_i = \sum_{i=1}^k \sum_{j=1}^{n_j} y_{ij} = \sum_{i=1}^4 \sum_{j=1}^{95} y_{ij} = 22820$	$\bar{Y} = \frac{\sum_{i=1}^k n_i \bar{y}_i}{n} = \frac{\sum_{i=1}^4 \bar{y}_i}{4} = 60.0526$

Matrix representation according to the model  $Y = X\beta + \varepsilon$

$$Y = \begin{bmatrix} 64 \\ 47 \\ \vdots \\ \vdots \\ 77 \\ 36 \end{bmatrix}_{(380 \times 1)} \quad X = \begin{bmatrix} 1 & 1 & 0 & 0 & \dots \\ 1 & \vdots & \vdots & \vdots & \vdots \\ \vdots & 0 & 1 & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & 0 & 1 & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & \vdots & \vdots & 0 & 1 \\ 1 & 0 & 0 & \dots & \vdots \end{bmatrix} \quad \beta = \begin{bmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_i \end{bmatrix}_{(i+1) \times 1} \quad \varepsilon = \begin{bmatrix} \varepsilon_{11} \\ \varepsilon_{12} \\ \vdots \\ \vdots \\ \varepsilon_{i1} \\ \varepsilon_{ij} \end{bmatrix}$$

is shaped.

$$\begin{matrix}
 & & \mu & \alpha_1 & \alpha_2 & \alpha_3 & \alpha_4 & & \\
 & & \left[ \begin{matrix} 1 & 1 & 0 & 0 & \dots \\ 1 & \vdots & \vdots & \vdots & \vdots \\ \vdots & 0 & 1 & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & 0 & 1 & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & \vdots & \vdots & 0 & 1 \\ 1 & 0 & 0 & \dots & \vdots \end{matrix} \right]_{380 \times 5} & & \left[ \begin{matrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_k \end{matrix} \right]_{5 \times 1} & + & \left[ \begin{matrix} \varepsilon_{11} \\ \varepsilon_{12} \\ \vdots \\ \varepsilon_{i1} \\ \varepsilon_{ij} \end{matrix} \right] \\
 Y = \left[ \begin{matrix} 64 \\ 47 \\ \vdots \\ \vdots \\ 77 \\ 36 \end{matrix} \right]_{(380 \times 1)} & = & & & & & & & \\
 Y & = & X & \beta & + & \varepsilon & & & 
 \end{matrix}$$

X: Design matrix with values 0 and 1 design matrix. Rank of the X design matrix is 4 and number of parameter is 5.  $X'$ ,  $X'X$  and  $X'Y$  matrix are created to determine the normal equation

$$X' = \left[ \begin{matrix} 1 & 1 & \dots & \dots & \dots & \dots & \dots & 1 \\ 1 & \vdots & 0 & \dots & \dots & \dots & \dots & \dots \\ 0 & \vdots & 1 & \dots & 0 & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \dots & 1 & \dots & 0 & \dots \\ \vdots & \vdots & \dots & \dots & \dots & \dots & 1 & \dots \end{matrix} \right]_{5 \times 380}$$

is obtained.

$$X'X = \left[ \begin{matrix} 1 & 1 & \dots & \dots & \dots & \dots & \dots & 1 \\ 1 & \vdots & 0 & \dots & \dots & \dots & \dots & \dots \\ 0 & \vdots & 1 & \dots & 0 & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \dots & 1 & \dots & 0 & \dots \\ \vdots & \vdots & \dots & \dots & \dots & \dots & 1 & \dots \end{matrix} \right] \left[ \begin{matrix} 1 & 1 & 0 & 0 & \dots \\ 1 & \vdots & \vdots & \vdots & \vdots \\ \vdots & 0 & 1 & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & 0 & 1 & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 1 & \vdots & \vdots & 0 & 1 \\ 1 & 0 & 0 & \dots & \vdots \end{matrix} \right] = \left[ \begin{matrix} 380 & 95 & 95 & 95 & 95 \\ 95 & 95 & 0 & 0 & 0 \\ 95 & 0 & 95 & 0 & 0 \\ 95 & 0 & 0 & 95 & 0 \\ 95 & 0 & 0 & 0 & 95 \end{matrix} \right]$$

as calculated.

$$X'Y = \left[ \begin{matrix} 1 & 1 & \dots & \dots & \dots & \dots & \dots & 1 \\ 1 & \vdots & 0 & \dots & \dots & \dots & \dots & \dots \\ 0 & \vdots & 1 & \dots & 0 & \dots & \dots & 0 \\ \vdots & \vdots & \vdots & \dots & 1 & \dots & 0 & \dots \\ \vdots & \vdots & \dots & \dots & \dots & \dots & 1 & \dots \end{matrix} \right] \left[ \begin{matrix} 64 \\ 47 \\ \vdots \\ \vdots \\ 77 \\ 36 \end{matrix} \right] = \left[ \begin{matrix} 22820 \\ 5398 \\ 5748 \\ 6001 \\ 5673 \end{matrix} \right]$$

as calculated.

$$(X'X)\hat{\beta} = X'Y$$

$$\begin{bmatrix} 380 & 95 & 95 & 95 & 95 \\ 95 & 95 & 0 & 0 & 0 \\ 95 & 0 & 95 & 0 & 0 \\ 95 & 0 & 0 & 95 & 0 \\ 95 & 0 & 0 & 0 & 95 \end{bmatrix} \begin{bmatrix} \hat{\mu} \\ \hat{\alpha}_1 \\ \hat{\alpha}_2 \\ \hat{\alpha}_3 \\ \hat{\alpha}_4 \end{bmatrix} = \begin{bmatrix} 22820 \\ 5398 \\ 5748 \\ 6001 \\ 5673 \end{bmatrix}$$

is found.  $\hat{\beta}$  vector system is estimated from this equation.  $X'X$  matrix is not full rank.  $\hat{\beta}$  vector is estimated using generalized inverse. Generalized G inverse of  $X'X$

$$G = \begin{bmatrix} 0 & 0 & 0 & 0 & 0 \\ 0 & 1/95 & 0 & 0 & 0 \\ 0 & 0 & 1/95 & 0 & 0 \\ 0 & 0 & 0 & 1/95 & 0 \\ 0 & 0 & 0 & 0 & 1/95 \end{bmatrix}$$

as calculated. Equality from  $\hat{\beta}^0 = GXY'$

$$\hat{\beta}^0 = \begin{bmatrix} 0 & 0 & 0 & 0 & 0 \\ 0 & 1/95 & 0 & 0 & 0 \\ 0 & 0 & 1/95 & 0 & 0 \\ 0 & 0 & 0 & 1/95 & 0 \\ 0 & 0 & 0 & 0 & 1/95 \end{bmatrix} \begin{bmatrix} 22820 \\ 5398 \\ 5748 \\ 6001 \\ 5673 \end{bmatrix}$$

$$= \begin{bmatrix} 0 \\ 5398/95 \\ 5748/95 \\ 6001/95 \\ 5673/95 \end{bmatrix} = \begin{bmatrix} 0 \\ 56.821 \\ 60.505 \\ 63.168 \\ 59.716 \end{bmatrix}$$

is obtained.

Sum of Squares Between Groups (SSB), Sum of Squares Within Groups (SSW) and General Sum of Squares (SST) are calculated to F test created.

$$SSB = \beta^{0'} X'Y - n\bar{y}^2$$

n = 380 (Number of student) and  $\bar{y} = 60,0526$  (average mark).

$$SSB = [0 \quad 56.821 \quad 60.505 \quad 63.168 \quad 59.168] \begin{bmatrix} 22820 \\ 5398 \\ 5748 \\ 6001 \\ 5673 \end{bmatrix} - 380(60.0526)^2$$

$$= 1372345.663 - 1370399.611 = 1946.663$$

is found or

$$SSB = \sum_{i=1}^k n_i (\bar{y}_i - \bar{y})^2 = \sum_{i=1}^4 n_i (\bar{y}_i - \bar{y})^2$$

$$= 95x(56,82 - 60,05)^2 + 95x(60,51 - 60,05)^2 + 95x(63,17 - 60,05)^2$$

$$+ 95x(59,72 - 60,05)^2 = 1946.663$$

as calculated.

Mean Square of Between Groups (MSB),

$$MSB = \frac{SSB}{k-1} = 1946.663 / 3 = 648.888$$

Mean Square Between Groups

$$MSB = Y'Y - \beta^{0'} X'Y$$

is shaped.

$$SSW = [64 \quad 47 \quad \dots \quad \dots \quad 77 \quad 36] \begin{bmatrix} 64 \\ 47 \\ \vdots \\ \vdots \\ 77 \\ 36 \end{bmatrix} - [0 \quad 56.821 \quad 60.505 \quad 63.168 \quad 59.716] \begin{bmatrix} 22820 \\ 5398 \\ 5748 \\ 6001 \\ 5673 \end{bmatrix}$$

$$= 1464054 - 1372345.663 = 91708.337$$

as calculated or

$$SSW = \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_j)^2 = \sum_{i=1}^4 \sum_{j=1}^{95} (y_{ij} - \bar{y}_j)^2$$

$$= (64 - 56,82)^2 + (47 - 56,82)^2 + \dots + (73 - 56,82)^2 + (67 - 60,51)^2 + \dots$$

$$+ (50 - 60,51)^2 + (69 - 63,17)^2 + \dots + (64 - 63,17)^2 + (42 - 59,72)^2 + \dots$$

$$+ (36 - 59,72)^2 = 91708.337$$

as found.

Mean Square Within Groups

$$MSW = \frac{SSW}{n-k} = \frac{91708.337}{380-4} = 243.905$$

**General Sum of Squares**

$$SST = Y'Y - ny^{-2} = 1464054 - 1370399.611 = 93654.389$$

as found.

F statistics,

$$F = \frac{SSB/(k-1)}{SSW/(N-k)} = \frac{MSB}{MSW} = \frac{648.888}{243.905} = 2.66$$

as obtained. Analysis of variance results the obtained are shown in Table 3.

**Table 3. Table for one-way classification**

Source of Variation	Sum of squares	Degrees of freedom	Mean square	F statistics
Between group	$\hat{\beta}^0 X'Y - ny^{-2} = 1946.052$	4-1=3	$MSB = 648.888$	$F = 2.66$
Within group	$Y'Y - \hat{\beta}^0 X'Y = 91708.337$	380-4=376	$MSW = 243.905$	
General	$Y'Y - ny^{-2} = 93654.389$	380-1=379		

The F statistics 2,66 is significant. Accordingly 3 and 376 degrees of freedom, the ANOVA table critical value is 2.60. Null hypothesis is reject. The difference among the achievement of students in a foreign language exam is statistically significant.

**Conclusion**

In this study,  $Y = X\beta + \varepsilon$  model is created.  $\beta^0$  vector as calculated

$$\beta^0 = [0 \ 56.821 \ 60.505 \ 63.168 \ 59.716]$$

SST, SSB and SSW values are found 93654.389, 1946.663 and 91708.337, respectively. MSB and MSW values are found 648.888 and 243.905, respectively. F statistics 2,66 is statistical significant. The system consisting of qualitative variable normal equations system is not inverse since it is not a full rank model. The solution was made with the generalized G inverse.

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